Using the New Products Margin
to Predict the Industry-Level Impact of Trade Reform

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Kehoe and Ruhl (2013) show that products that are traded very little or not at all account disproportionately for aggregate changes in bilateral trade following trade liberalization or rapid growth experiences, but not over the business cycle.

Hypothesis: Industries with more trade due to these little-traded and nontraded products should experience more growth following trade liberalization.

Product: A 5-digit SITC, rev. 2 code. There are 1,836 products.

Industry: A 3-digit ISIC code. There are 38 industries. (We are only interested in industries that produce goods in merchandise trade agriculture, mining and extraction, and manufacturing.)

We map products into industries using concordance developed by Muendler (2009).

Notice that each industry, on average, consists of 48.3 products.

111 Agriculture and livestock production
113 Hunting, trapping and game propagation
121 Forestry
122 Logging
130 Fishing
210 Coal mining
220 Crude petroleum and natural gas production
230 Metal ore mining
290 Other mining
311-312 Food manufacturing
313 Beverage industries
314 Tobacco manufactures
321 Manufacture of textiles
322 Manufacture of wearing apparel, except footwear
323 Manufacture of leather and products of leather, leather substitutes and fur
324 Manufacture of footwear
331 Manufacture of wood and wood and cork products, except furniture
332 Manufacture of furniture and fixtures, except primarily of metal
341 Manufacture of paper and paper products

342 Printing, publishing and allied industries
351 Manufacture of industrial chemicals
352 Manufacture of other chemical products
353
354

390 Other manufacturing industries

## The New Product, or Extensive, Margin

We sort each of the 1,836 products by average amount of trade over the first three years of our period

We then place each product into bins sequentially until each bin accounts for 10 percent of total trade in the base period.

We define Least Traded Products (LTP) to be the products in the final 10 percent bin, the products with the least amount of trade over the first three years.

Composition of Exports: Canada to United States 1988-2009


Composition of Exports: Spain to Germany 1978-1987


Composition of Exports: Spain to Germany 1988-2008


Composition of Exports: Germany to Spain 1978-1987


Composition of Exports: Germany to Spain 1988-2008


## Comparison to other extensive margins

Most of the literature uses a fixed cutoff when deciding whether a product is part of the extensive margin, Feenstra (1994) uses a value of $\$ 0$, and Evenett and Venables (2002) use $\$ 50,000$.

In contrast, our measure varies by country pairs. The cutoff for EcuadorPeru differs from the cutoff for U.S.-Canada.

We keep our set of extensive margin products fixed, as opposed to focusing on movement into and out of the extensive margin.

## Predicting changes in trade by industry

Compute the fraction of trade in each industry accounted for by LTP $s_{j}$ in the base period $t_{0}$. Predict

$$
\begin{gathered}
z_{j}=\alpha+\beta s_{j} \\
z_{j}=\frac{X_{j i t}^{k} / G D P_{i t}}{X_{j i_{0}}^{k} / G D P_{i t_{0}}}-1
\end{gathered}
$$

and $X_{j i t}^{k}$ are exports of industry $j$ from country $i$ to country $k$ in year $t$. We use experience from previous trade reforms (in this case NAFTA) to estimate $\alpha$ and $\beta$. Our hypothesis is that $\beta>0$.

Least Traded Products predictions compared to those of Yaylaci-Shikher (forthcoming)

|  | Korea to United States |  |  | United States to Korea |  |  |
| :--- | ---: | ---: | ---: | ---: | ---: | ---: |
| industry | Yaylaci- <br> Shikher <br> predictions | LTP <br> predictions | $\mathbf{2 0 0 5}$ <br> fraction <br> LTP | Yaylaci- <br> Shikher <br> predictions | LTP <br> predictions | 2005 <br> fraction <br> LTP |
| Chemicals | 28.2 | 54.00 | 0.36 | 30.3 | 20.70 | 0.16 |
| Electrical mach. | 15.5 | -0.44 | 0.02 | 41.0 | -3.02 | 0.04 |
| Food | 70.1 | 86.03 | 0.56 | 422.3 | 26.63 | 0.19 |
| Other machinery | 8.9 | 9.17 | 0.08 | 31.9 | 6.86 | 0.09 |
| Medical | 9.9 | 74.82 | 0.49 | 45.0 | -1.05 | 0.05 |
| Metals | 9.3 | 17.18 | 0.13 | 17.0 | 28.61 | 0.20 |
| Nonmetals | 20.5 | 39.59 | 0.27 | 38.7 | 80.00 | 0.46 |
| Other | 11.8 | 50.80 | 0.34 | 28.5 | 40.47 | 0.26 |
| Paper | 1.4 | 105.24 | 0.68 | 5.5 | 6.86 | 0.09 |
| Petroleum | 2.2 | 15.57 | 0.12 | 7.2 | -5.00 | 0.03 |
| Metal products | 14.2 | 62.01 | 0.41 | 33.8 | 20.70 | 0.16 |
| Rubber | 19.8 | 10.77 | 0.09 | 48.0 | 22.68 | 0.17 |
| Textile | 56.3 | 58.81 | 0.39 | 63.5 | 117.56 | 0.65 |
| Transport. equip. | 23.3 | -2.04 | 0.01 | 33.9 | -5.00 | 0.03 |
| Wood | 7.9 | 29.99 | 0.21 | 21.1 | 38.49 | 0.25 |
| Chemicals | 28.2 | 54.00 | 0.36 | 30.3 | 20.70 | 0.16 |
| KS-LTP weighted correlation |  | $\mathbf{0 . 4 3}$ |  |  | $\mathbf{0 . 1 9}$ |  |

Kehoe (2005) showed that several of the leading models built to predict the industry level effects of NAFTA performed poorly

We confirm this finding for Brown-Deardorff-Stern (BDS), Cox-Harris, and Sobarzo models over the 1989-2009 period.

Focus on the BDS model since it has bilateral trade predictions for all importer-exporter pairs between Canada, Mexico, and the U.S.

## Methodology for evaluating the NAFTA models

We compute the weighted correlation coefficient between the model predictions and the results from the data

We also compute the weighted regression coefficients $a$ and $b$ from

$$
\min _{a, b} \sum_{j=1}^{23} \omega_{j}\left(a+b z_{j}^{\text {model }}-z_{j}^{\text {data }}\right)^{2}
$$

Here $a$ indicates how well the models did in matching average change ( $a=0$ is ideal) and $b$ indicates how well the models did in matching the signs and magnitudes of the changes ( $b=1$ is ideal)

Changes in Canada-U.S. trade relative to exporter's GDP (percent)

|  | Canada to U.S. |  |  | U.S. to Canada |  |  |
| :--- | ---: | ---: | ---: | ---: | ---: | ---: |
| industry | 1989-2009 <br> data | BDS <br> model | $\mathbf{1 9 8 9}$ <br> fraction <br> LTP | 1989-2009 <br> data | BDS <br> model | $\mathbf{1 9 8 9}$ <br> fraction <br> LTP |
| Agriculture | 12.5 | 3.4 | 0.26 | -6.4 | 5.1 | 0.19 |
| Mining and quarrying | 237.6 | 0.4 | 0.05 | 51.3 | 1.0 | 0.16 |
| Food | 101.2 | 8.9 | 0.24 | 124.1 | 12.7 | 0.25 |
| Textiles | 42.4 | 15.3 | 0.77 | -35.9 | 44.0 | 0.52 |
| Clothing | 50.2 | 45.3 | 0.59 | -3.0 | 56.7 | 1.00 |
| Leather products | -67.7 | 11.3 | 1.00 | -64.0 | 7.9 | 0.61 |
| Footwear | -49.9 | 28.3 | 1.00 | -67.2 | 45.7 | 0.34 |
| Wood products | -54.5 | 0.1 | 0.01 | -30.6 | 6.7 | 0.07 |
| Furniture and fixtures | -46.6 | 12.5 | 0.00 | 22.5 | 35.6 | 0.00 |
| Paper products | -65.9 | -1.8 | 0.04 | 13.7 | 18.9 | 0.15 |
| Printing and publishing | 0.7 | -1.6 | 0.12 | -19.6 | 3.9 | 0.05 |
| Rubber products | 45.8 | 9.5 | 0.10 | 30.2 | 19.1 | 0.05 |
| Chemicals | 99.6 | -3.1 | 0.38 | 50.2 | 21.8 | 0.24 |
| Petroleum products | -79.8 | 0.5 | 0.07 | -43.1 | 0.8 | 0.13 |
| Glass products | -45.7 | 30.4 | 0.40 | -20.0 | 4.4 | 0.23 |
| Nonmetal mineral products | -0.4 | 1.2 | 0.38 | -1.9 | 11.9 | 0.59 |
| Iron and steel | -12.7 | 12.9 | 0.36 | 53.5 | 11.6 | 0.28 |


| Nonferrous metals | -20.9 | 18.5 | 0.07 | -20.8 | -6.7 | 0.11 |
| :--- | ---: | ---: | ---: | ---: | ---: | ---: |
| Metal products | 17.7 | 15.2 | 0.20 | -5.3 | 18.2 | 0.16 |
| Nonelectrical machinery | -8.4 | 3.3 | 0.21 | -38.9 | 9.9 | 0.08 |
| Electrical machinery | -16.4 | 14.5 | 0.15 | -42.6 | 14.9 | 0.05 |
| Transportation equipment | -44.3 | 10.7 | 0.01 | -37.8 | -4.6 | 0.01 |
| Misc. manufactures | 56.1 | -2.1 | 0.45 | -19.2 | 11.5 | 0.15 |
| weighted corr. with data |  | -0.28 | 0.30 |  | 0.39 | 0.54 |
| regression coeff. $\boldsymbol{a} \backslash \boldsymbol{\alpha}$ |  | 21.82 | -20.42 |  | -26.62 | -34.54 |
| regression coeff. $\boldsymbol{b} \backslash \boldsymbol{\beta}$ | -3.33 | 185.24 |  | 1.34 | 175.84 |  |
| BDS-LTP weighted corr. |  |  | -0.11 |  |  | 0.70 |

Results for the BDS model: the BDS model fared poorly in predicting industry level changes in bilateral trade

| exporter | importer | correlation | $\boldsymbol{a}$ | $\boldsymbol{b}$ |
| :--- | :--- | ---: | ---: | ---: |
| Canada | Mexico | -0.10 | 645.29 | -7.94 |
| Canada | United States | -0.28 | 21.82 | -3.33 |
| Mexico | Canada | 0.06 | 135.79 | 0.16 |
| Mexico | United States | -0.13 | 66.64 | -0.11 |
| United States | Canada | 0.39 | -26.62 | 1.34 |
| United States | Mexico | -0.06 | 88.47 | -0.24 |
| weighted average | -0.00 | 19.83 | -0.94 |  |
| pooled regression | 0.06 | 10.54 | 0.17 |  |

Correlation is the weighted correlation of predictions with the data.

Results for the LTP exercise: the LTP exercise fares much better in predicting industry level changes in bilateral trade

| exporter | importer | correlation | $\boldsymbol{\alpha}$ | $\boldsymbol{\beta}$ |
| :--- | :--- | ---: | ---: | ---: |
| Canada | Mexico | 0.55 | 254.23 | 4468.37 |
| Canada | United States | 0.30 | -20.42 | 185.24 |
| Mexico | Canada | 0.33 | 115.16 | 286.39 |
| Mexico | United States | 0.19 | 51.52 | 77.54 |
| United States | Canada | 0.54 | -34.54 | 175.84 |
| United States | Mexico | 0.47 | 62.31 | 265.44 |
| weighted average | 0.39 | -5.74 | 87.29 |  |
| pooled regression | 0.24 | -5.30 | 181.18 |  |

Comparison of the BDS results and LTP exercise results: LTP
exercise performs better the BDS model for every country pair.

|  |  | BDS |  |
| :--- | :--- | ---: | ---: |
| exporter | importer | correlation correlation |  |
| Canada | Mexico | -0.10 | 0.55 |
| Canada | United States | -0.28 | 0.30 |
| Mexico | Canada | 0.06 | 0.33 |
| Mexico | United States | -0.13 | 0.19 |
| United States | Canada | 0.39 | 0.54 |
| United States | Mexico | -0.06 | 0.47 |
| weighted average | -0.00 | 0.39 |  |
| pooled regression | 0.06 | 0.24 |  |

## Growth in data versus BDS predicted growth

Canadian exports to the United States



Our exercise shows that looking at the share of least traded products in an industry is a useful predictor of which industries will experience the most growth following trade liberalization.

Major downside to our method: As of now it is atheoretical.

It is our hope that our results will spur the development of models able to account for the importance of the new product margin in trade.

## Robustness:

The $\alpha$ and $\beta$ computed from our industry-level regressions tell us how much LTP and non-LTP products grew on average

We compare these industry-based estimates to the average growth rates computed directly from the product data.

The industry level growth rates will not account for products with zero trade in 1989, while the product level growth rates will. If the estimated growth rates are similar, it indicates the important products are the ones with small, but positive trade.

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We find a weighted correlation of 0.97 for $\alpha$ and 0.91 for $\beta$

Changes in North American trade relative to exporter's GDP:
Estimates from industry data versus estimates from product data

|  |  |  | industry data |  | product data |  |
| :--- | :--- | :--- | ---: | ---: | ---: | ---: |
| exporter | importer | period | $\boldsymbol{\alpha}$ | $\boldsymbol{\beta}$ | $\tilde{\boldsymbol{\alpha}}$ | $\tilde{\boldsymbol{\beta}}$ |
| Canada | Mexico | $89-09$ | 273.01 | 4253.33 | 452.67 | 2483.99 |
| Canada | United States | $89-09$ | -16.89 | 149.92 | -14.57 | 126.73 |
| Mexico | Canada | $89-09$ | 107.47 | 363.23 | 96.13 | 476.67 |
| Mexico | United States | $89-09$ | 54.92 | 43.54 | 46.89 | 123.86 |
| United States | Canada | $89-09$ | -28.22 | 112.55 | -21.61 | 46.48 |
| United States | Mexico | $89-09$ | 65.96 | 228.93 | 78.46 | 103.92 |
| weighted correlation $\boldsymbol{\alpha}, \tilde{\boldsymbol{\alpha}}$ |  |  |  |  |  |  |
|  | 0.97 |  |  |  |  |  |
| weighted correlation $\boldsymbol{\beta}, \tilde{\boldsymbol{\beta}}$ |  |  | 0.91 |  |  |  |

C. Arkolakis (2010), "Market Penetration Costs and the New Consumers Margin in International Trade."



## Robustness

We also find that our results hold when changing our end dates. For example if we use 1988(when available)-2007 to avoid the great recession.

We also find that, for goods for which we have both price and quantity data, after deflating by the exporter's PPI - most changes in value are driven by changes in quantity.

Our exercise similarly performs well when compared to alternative models used to predict the effects of NAFTA, for example Cox-Harris for Mexico and Sobarzo for Canada.

Changes in North American trade deflated by Exporter's PPI: Growth due to quantities versus change due to prices

|  |  | average share of |  |
| :--- | :--- | ---: | ---: |
| total growth |  |  |  |


| sector | Exports to North America |  |  | Imports from North America |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | $\begin{gathered} \text { 1989- } \\ 2009 \\ \text { data } \\ \hline \end{gathered}$ | Sobarzo growth rate | $\begin{gathered} 1989 \\ \text { fraction } \\ \text { least } \\ \text { traded } \\ \hline \end{gathered}$ | $\begin{gathered} \text { 1989- } \\ 2009 \\ \text { data } \\ \hline \end{gathered}$ | Sobarzo growth rate | $\begin{gathered} 1989 \\ \text { fraction } \\ \text { least } \\ \text { traded } \end{gathered}$ |
| Agriculture | -15.3 | -11.1 | 0.07 | 3.2 | 3.4 | 0.10 |
| Beverages | 161.8 | 5.2 | 0.01 | 85.2 | -1.8 | 0.32 |
| Chemicals | 34.1 | -4.4 | 0.60 | 104.2 | -2.7 | 0.23 |
| Electrical Machinery | 54.7 | 1.0 | 0.02 | 6.6 | 9.6 | 0.01 |
| Food | 100.8 | -6.9 | 0.41 | 46.7 | -5.0 | 0.15 |
| Iron and Steel | 19.6 | -4.9 | 0.37 | 23.1 | 17.7 | 0.24 |
| Leather | -64.6 | 12.4 | 0.53 | 2.5 | -0.4 | 0.67 |
| Metal Products | 86.2 | -4.4 | 0.30 | 24.8 | 9.5 | 0.14 |
| Mining | 27.7 | -17.0 | 0.01 | 15.0 | 13.2 | 0.17 |
| Nonelectrical Machinery | 166.5 | -7.4 | 0.12 | 38.3 | 20.7 | 0.09 |
| Nonferrous Metals | 36.8 | -9.8 | 0.13 | 37.1 | 9.8 | 0.10 |
| Nonmetallic Min. Prod. | -16.0 | -6.2 | 0.26 | 5.3 | 10.9 | 0.49 |
| Other Manufactures | 88.4 | -4.5 | 0.23 | 26.1 | 4.2 | 0.16 |
| Paper | -35.9 | -7.9 | 0.30 | -4.1 | -4.7 | 0.07 |
| Petroleum | -98.0 | -19.5 | 0.12 | -81.6 | -6.8 | 0.06 |
| Rubber | 158.9 | 12.8 | 0.43 | 78.3 | -0.1 | 0.06 |
| Textiles | 69.5 | 1.9 | 0.76 | 48.3 | -1.2 | 0.44 |
| Tobacco | -61.3 | 2.8 | 1.00 | 333.0 | -11.6 | 1.00 |
| Transportation Equip. | 126.1 | -5.0 | 0.02 | 26.7 | 11.2 | 0.02 |
| Wearing Apparel | 197.2 | 30.0 | 0.23 | -17.2 | 4.5 | 0.20 |
| Wood | 30.8 | -8.5 | 0.04 | -34.0 | 11.7 | 0.05 |
| weighted correlation with data |  | 0.43 | 0.02 |  | -0.12 | 0.47 |
| regression coefficient $a \backslash \alpha$ |  | 62.91 | 81.13 |  | 30.91 | 9.61 |
| regression coefficient $\boldsymbol{b} \backslash \boldsymbol{\beta}$ |  | 7.92 | 3.06 |  | -0.49 | 175.76 |

Changes in Canadian trade relative to Canadian GDP in the Cox-Harris Model (Percent)

| sector | Exports to World |  |  | Imports from World |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | $\begin{gathered} \text { 1989- } \\ 2009 \\ \text { data } \end{gathered}$ | C-H growth rate | $\begin{gathered} 1989 \\ \text { fraction } \\ \text { least } \\ \text { traded } \\ \hline \end{gathered}$ | $\begin{gathered} \text { 1989- } \\ 2009 \\ \text { data } \\ \hline \end{gathered}$ | C-H growth rate | $\begin{gathered} 1989 \\ \text { fraction } \\ \text { least } \\ \text { traded } \\ \hline \end{gathered}$ |
| Agriculture | 39.1 | -4.1 | 0.13 | 13.4 | 7.2 | 0.18 |
| Chem. \& Misc. Man. | 70.9 | 28.1 | 0.34 | 59.1 | 10.4 | 0.20 |
| Fishing | -30.9 | -5.4 | 0.05 | 32.9 | 9.5 | 0.22 |
| Food, Bev., and Tobacco | 95.5 | 18.6 | 0.22 | 86.6 | 3.8 | 0.19 |
| Forestry | -24.8 | -11.5 | 0.15 | 4.5 | 7.1 | 0.24 |
| Machinery and Appl. | 11.7 | 57.1 | 0.19 | -6.6 | 13.3 | 0.06 |
| Mining | 117.0 | -7.0 | 0.03 | 103.0 | 4.0 | 0.06 |
| Nonmetallic Minerals | 20.9 | 31.8 | 0.64 | 3.4 | 7.3 | 0.32 |
| Refineries | -67.8 | -2.7 | 0.06 | -71.9 | 1.5 | 0.03 |
| Rubber and Plastics | 107.3 | 24.5 | 0.22 | 56.0 | 13.8 | 0.07 |
| Steel and Metal Products | 6.6 | 19.5 | 0.15 | 33.2 | 10.0 | 0.17 |
| Textiles and Leather | 18.4 | 108.8 | 0.86 | -1.9 | 18.2 | 0.33 |
| Transportation Equip. | -37.5 | 3.5 | 0.01 | -19.7 | 3.0 | 0.01 |
| Wood and Paper | -58.5 | 7.3 | 0.02 | 12.8 | 7.2 | 0.09 |
| weighted correlation with data |  | 0.06 | 0.40 |  | 0.04 | 0.48 |
| regression coefficient $\boldsymbol{a} \backslash \boldsymbol{\alpha}$ |  | 2.00 | -13.73 |  | 9.77 | -7.55 |
| regression coefficient $\boldsymbol{b} \backslash \boldsymbol{\beta}$ |  | 0.16 | 199.46 |  | 0.30 | 199.46 |

